PART 2: AUTOREGRESSIVE SERIES

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1. INTRODUCTION

In his classical paper of 1927, where the autoregressive scheme was introduced, Yule developed a method, for fitting the scheme to a given time series. In later contributions by Wold (1938), Kendall (1947), Quenouille (1947) and others, the problems of fitting and parameter estimation were given further consideration. In the present paper an attempt is made with the help of model samples to see how far the methods given by these authors hold good when the length of the time series is small. In this part, models conforming to stationary time series of autoregressive type have been considered. Quenouille (1947) has given a test for testing the goodness of fit of an a priori known model to an observational time series when the length of the series is large. The suitability or otherwise of this test for time series of small length is considered in this paper. He has also given a method of fitting an autoregressive model and to test for the fitted model. In this paper this method is applied to time series of small length and alternative methods of fitting autoregressive models are considered. The possibility of fitting a suitable moving average model to a stationary time series of autoregressive type is also considered.

These investigations were undertaken primarily to get material for further studies, and to gain instructive insight into the problems at issue than to arrive at definite conclusions. Nevertheless, the conclusions drawn here are pointers to further studies.

2. Models for study

2.0. Construction of models: The autoregressive models chosen for study are:

Model I
$$\xi_i + \sigma \xi_{i-1} = \eta_i$$
 ... (1)

Model II
$$\xi_1 + a_1 \xi_{1-1} + a_2 \xi_{1-2} = \eta_1$$
 ... (2)

where & is the variable at time point t,

$$a = -.8$$
; $\sigma_1 = -.7$; $a_2 = .6125$;

and 7, is a random normal variable with

$$E(\eta_i) = 0$$
, $E(\eta_i^t) = 1$ and $E(\eta_i \eta_{i'}) = 0$ for $t \neq t'$.

These models were constructed by using random normal deviates given in Tracts for Computers No. XXV by Herman Wold. Twenty five samples of length 35 and fifty samples of length 15 were constructed for both the models. In building up the models it was necessary to know ξ_1 for model I and ξ_1 and ξ_2 for model II with the help of which other values of ξ_1 for $t \geq 2$ could be built up from equations (1) and (2).

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f, can also be written in the form

$$\xi_1 = \eta_1 + b_1 \eta_{1-1} + b_2 \eta_{1-2} + \dots$$
 (3)

where by is given by

Model I
$$ab_{i-1} + b_i = 0$$

Model II $a_rb_{i-1} + a_rb_{i-1} + b_i = 0$

where $b_i = 0$ for i < 0, and $b_a = 1$.

The initial values of ξ , were built up from equation (3). Values of ξ , for $t \ge 2$ were then built up from equations (1) and (2).

Values of random normal deviates given in pages 4, 5, 6, 7 and 8 for model I and pages 7, 14 and 21 for model II were used for calculating the initial values of ξ_1 in the model samples. These satisfy almost all tests given in the introduction to Tracts for Computers No. X.X.V.

2.1. Serial correlation coefficients: Yulo (1927) has shown that the serial correlation coefficients for the a priori known model, say ρ_s , should satisfy the recursive relation

Model I
$$\rho_{a} + a\rho_{a-1} = 0$$
 ... (4)

Model II
$$\rho_0 + \sigma_1 \rho_{n-1} + \sigma_2 \rho_{n-2} = 0$$
 ... (5)

when $\rho_0 = 1$ and $\rho_{-1} = \rho_1$.

But Wold (1938) showed that equations (4) and (5) are valid only for s>0; for $s\leqslant 0$, the zero in the right hand side of the equations should be replaced

by
$$\frac{\delta_k}{\Sigma b_k^{-2}}$$
. The equations (4) and (5) may therefore be called Yule-Wold relations.

The values of ρ_s for s equal to 1 to 12 in the case of samples of length 35 and s equal to 1 to 8 in the case of samples of length 15 were worked out with the help of these recursive relations. The value of ρ_s are shown in Tables 1.1 to 1.4 at the end of this Part.

The corresponding serial correlation coefficients for the samples, say r_s , where r_s represents the product moment correlation coefficient between ξ_1 and ξ_{1-s} , were worked out with the help of Hollerith machines. The values of r_s are shown in Tables 1.1 to 1.4.

2.2. Bias in the serial correlation coefficients from small samples: Since the values of r, were calculated from the series of length 35 and 15 only, these are biased estimates of p.. The bias can be calculated by working out the expectation of r, which is

$$E(\mathbf{r}_{i}) = E\left\{\frac{\sum_{i=1}^{T_{i}} \xi_{i} \xi_{i+1} - \left(\sum_{i=1}^{T_{i}} \xi_{i} \sum_{i=1}^{T_{i}} \xi_{i+1}\right) / (T-S)}{\sqrt{D_{1}^{T_{i}} D_{2}^{T_{i}}}}\right\} \qquad ... \quad (6)$$

where

$$D_{i}^{t} = \sum_{i=1}^{T-s} \xi_{i}^{t} - \left(\sum_{i=1}^{T-s} \xi_{i}\right)^{s} / (T-S)$$

$$D_{i}^{t} = \sum_{i=1}^{T} \xi_{i}^{t} - \left(\sum_{i=1}^{T} \xi_{i}\right)^{s} / (T-S)$$

and T=length of the time series. From the above,

$$E(r_s) \sim \frac{E\left(\sum_{i=1}^{r_{s+1}} \xi_i \xi_{i+s}\right) - E\left\{\left(\sum_{i=1}^{r_{s+1}} \xi_i \sum_{i=1}^{r_{s+1}} \xi_{i+s}\right) / (T-S)\right\}}{E(\sqrt{D_i} \cdot D_i^{r_s})}$$
 ... (7)

When T is fairly large D_1^* can be taken as approximately equal to D_2^* .

It can be shown that

$$E(D_1^{\bullet}) \sim E(D_2^{\bullet}) = \sigma^{\circ} \left[(T-S) - \frac{1}{(T-S)} \left\{ (T-S) + 2(T-S-1)\rho_1 + \dots 2\rho_{T-1-1} \right\} \right] \dots$$
 (8)

where $\sigma^1 = E(\eta^1) = 1$.

The numerator in equation (7) can be written as C-D where

$$C = (T - S)\rho_{\bullet} \qquad ... \quad (9)$$

and

$$D = \frac{E - F}{T - S}$$
 where

$$E = \left\{\sum_{i=1}^{l+t-1} (i + \overline{i-t} - \overline{t-s})\rho_i\right\} + i \quad \text{where} \quad i = T - S \quad \dots \quad (10)$$

$$F = \left\{ \sum_{i=1}^{i_1 \cdot s_1 - 1} (i_1 + \overline{i_1 - t} - \overline{t - s_1}) \rho_i \right\} + i_1 \text{ where } i_1 = S, \text{ and } S_1 = T - 2S \quad \dots \quad (11)$$

and (i-t)=0 for $i \le t$; $(i_1-t)=0$ for $i_1 \le t$; t-s=0 for $t \le s$; $(t-s_1)=0$ for $t \le s_1$.

Values of $E(r_*)$ worked out by the above formulae are given in Tables 1.1 to 1.4. Comparing $E(r_*)$ with ρ_* it is seen that the bias is of a small order in the case of the models considered in this paper.

The average correlograms are shown in Figs. 1.1 to 1.4. It is seen that the observed correlograms follow the theoretical correlograms fairly closely.

3. TEST FOR THE GOODNESS OF FIT OF A PRIORI KNOWN MODELS

Quenouille (1947) has given a test criterion for testing the goodness of fit of a priori known models to time series of large length. His test is to build up R_* given by

Model I $R_s = r_s + A_1 r_{s-1} + A_2 r_{s-2}$... (12) where $A_1 = 2a$: $A_2 = a^2$.

a being known and equal to -0.8 in our case:

Model II $R_s = r_s + A_1 r_{s-1} + A_2 r_{s-2} + A_3 r_{s-2} + A_4 r_{s-4}$... (13) where $A_1 = 2a_1; A_2 = a_1^3 + 2a_2; A_3 = 2a_1a_2; A_4 = a_2^3,$

 a_1 and a_2 being known and equal to -0.7 and 0.6125 respectively.

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Quenouille has shown that R, are independently and normally distributed about zero with variance approximately equal to $(\frac{1-a^2}{a^2-b^2})^2$ in the case of model (I) and to

$$\frac{1}{2!-S} \left[\frac{(1-a_1)^{2}(1+a_2)^{2}-a_1^{2}}{(1+a_1)} \right]^{2}$$

in the case of model II, and hence $\frac{R_s^2}{VR_s^2}$ are distributed like χ^2 with 1 degree of freedom. Further he has shown that $(r_1-\rho_1)$ is distributed normally with zero mean and variance equal to $\frac{1-\rho^2}{T}$ in the case of model I and hence $\frac{(r_1-\rho_1)^2}{V(r_1)}$ is distributed like x2 with 1 degree of freedom. In the case of model II it was assumed that $(r_1-\rho_1)$ and $(r_2-\rho_2)$ are distributed normally with zero mean and variance approximately equal to $\frac{1-\rho_1^2}{T}$ and $\frac{1-\rho_2^1}{T}$.

The values of x2 were worked out by these formulae. The frequency distribution of x3 with 1 degree of freedom and also of total x2 for R4 are shown in Table 2.1. It is seen that a large percentage of x2 are significant at 5% level and that the observed frequency distributions of x2 are not in good agreement with the theoretical distribution of x1. The observed frequencies in different class intervals are smaller than the corresponding theoretical frequencies upto $P(\chi^2)$ equal to 0.50 and larger beyond.

Quenouille's test for goodness of fit of a priori known models was applied to the average serial correlations with and without correction for bias and the values of χ^2 obtained are as shown in Tables 2.2 and 2.3. It is seen that only 3 out of the 16 χ^2 's with correction for bias in r. are significant at 5% level in the case of model II.

Since there was a preponderence of x2 significant at 5% level, an attempt was made to see if significant x2's correspond to larger values of S. Table (2.4) shows the number of x2 significant at 5% level for different values of S. It is seen that there is no such tendency. TABLE 2.4. No. of X WITH | D.F. SIGNIFICANT AT 5% LEVEL

lag (#)	ξ ₁ -0.7 ξ ₁₋₁ +6	
	T=35	T=15
(1)	(2)	(3)
1	5	12
1 2 3 4 5 6 7 8	5 4 2 1 2 1 2 0 5	9
3	2	Ð
4	1	7
5	2	5
6	1	14
7	2	
8	0	
D	8	
10	**	••
11		
12		••
otal	22	81

4. FITTING OF MODELS

4.0. Fitting of autoregressive models to time series: In fitting autoregressive models to time series, the method adopted by Yulo (1927), Wold (1938), and Kendall (1947) is to equate the first serial correlation r₁ to ρ₁ in the model I and to equate r₁ and r₂ to ρ₁ and ρ₂ in the model II. For model II, a₁ and a₂ can be calculated from the relation

$$\rho_1 + a_1 + a_1 \rho_1 = 0 \qquad ... \quad (14)$$

$$\rho_1 + \tilde{a}_1 \rho_1 + a_2 = 0 \qquad ... \tag{15}$$

where ρ_1 and ρ_2 are equated to r_1 and r_2 .

This method was adopted for model II and the values of a_1 and a_2 for the twenty five samples of length 35 are shown in Table 3. The absolute percentage differences of these estimated values from the theoretical values range from 1.2% to 75.4%.

Kendall (1947) suggested that a, and a can be obtained by minimising

$$\sum_{n=1}^{\infty} (r_n + a_1 r_{n-1} + a_2 r_{n-2})^2$$

The values of σ_1 and σ_2 obtained by this method are shown in Table 3. The absolute percentage difference of these estimated values from theoretical values range from 1.4% to 61.1% and are generally less than in the case of earlier method of fitting.

A third alternative method was also tried. The values of A., A., A. and A.

were determined by minimising $\sum_{i=1}^{12} R_i t^i$ where

$$R_{a} = r_{a} + A_{1}r_{a-1} + A_{2}r_{a-2} + A_{3}r_{a-3} + A_{4}r_{a-4}. \qquad ... (10)$$

This method suggests itself if we consider Quenouille's test.

Having got A_1 , A_2 , A_3 and A_4 the best value of a_1 and a_2 were estimated by the following procedure. These constants can be written as

$$A_1 = 2a_1 = 2a'_1 + 2\delta a_1,$$
 ... (17)

$$A_1 = a_1^1 + 2a_1 = a_1'^1 + 2a_1'\delta a_1 + 2a_1' + 2\delta a_1,$$
 ... (18)

$$A_{2} = 2a_{1}a_{2} = 2a'_{1}a'_{2} + 2a'_{1}\delta a_{1} + 2a'_{2}\delta a_{1}, \qquad ... (19)$$

$$A_4 = a_2^2 = a_2^{\prime 2} + 2\delta a_2, \qquad ... \tag{20}$$

where $a_1 = a'_1 + \delta a_1$ and $a_2 = a'_2 + \delta a_2$, and second and higher powers and products of δ 's are omitted as negligible. a'_1 and a'_2 are first approximations to a_1 and a_2 and these are taken from the second method of fitting. The values of a_1 and a_2 were worked out for 5 samples only by this method. These values are shown in Table 3 and they differ very much from the theoretical values.

4.1. Test for goalness of fit of fitted authors/ressive models: Quenouillo's test of goodness of fit for fitted values of a₁ and a₂ by the first two methods was carried out and the distribution of x² with 1 d.f. and of total χ² with 10 d.f. are shown in

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- Table 4. The observed frequencies are more than the theoretical frequencies upto $P(\chi^2)$ equal to 0.60 and less beyond.

4.2. Fitting of moving average models to stationary time series of autoregressive type: The method of fitting moving average schemes to time series data is described in Part 1. The same method was adopted in fitting a three-constant moving average scheme to a model time series of autoregressive type with two constants. Herman Wold (1949) has given a large sample test for the goodness of fit of moving averages to time series data. This test was adopted for testing the goodness of fit. The study was carried out with 6 samples only and the values of χ^2 obtained with 1 d.f. are shown in Table 5. It is seen that none out of 54 have turned out to be significant. This indicates that the moving average scheme with three constants fits an autoregressive time series with two constants fairly well. It may be that a moving average scheme with more number of constants fairly well.

5. SUMMARY OF CONCLUSIONS

- (i) Quenouillo's test of goodness of fit of a priori known models to time series data is not very satisfactory when the length of the time series is short.
- (ii) Kendall's method of fitting autoregressive models with h constants to time series data by minimising

$$\sum_{n=1}^{\infty} (r_0 + a_1 r_{s-1} + ... + a_b r_{s-b})^2$$

appears to be better than that by equating $r_1, r_2, ..., r_k$ to $\rho_1, \rho_2, \rho_3, ..., \rho_k$ respectively.

(iii) Moving average schemes with more than h constants appear to fit the autoregressive time series with h constants fairly well.

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Table 1.1 Values of skrial correlation coefficients $\{r_n\}$ Mookl 1—Automegressive Model: $\{r_n=0.8, \{r_n\}, \neg q_n\}$

T-15

sample no.	71	7,	r 3	r.	r,	·.	F1	r,
1	3	3	4	5	0	7	8	9
ı	.6083	.1271	3517	 .7789	6276	.1214	.5789	.7065
2	.6053	.0741	0416	3508	6730	-0072	6076	6540
3	.7420	.7209	.4992	.6444	.4231	1260.	.2123	3822
Ĭ.	.4612	. 1020	2140	2891	5204	4390	.0113	.7618
5	.8012	.4819	.0634	4283	7524	8011	6593	3403
6	.3931	2709	5635	4330	.0130	.2903	.4194	-4567
7	.6561	.1980	3007	4333	4877	-,6370	6739	9473
8	.7239	.7399	.4147	.0764	.1590	6194	~.3686	6633
9	.8330	. 6266	.3022	0909	5419	792t	7024	9556
10	.3715	.1110	2367	6151	5540	7178	.1049	.4350
11	.3276	1610	1748	3857	3910	1359	.2844	.0769
12	, UKGO	.5200	.4563	.5146	, 2362	1684	4237	1706
13	.2353	,4352	.0143	0135	2091	291s	3440	2612
14	.8264	.7881	.7785	.7075	.7260	.3598	.6000	.4560
is	.3825	.0489	1031	.0493	4162	6808	020G	.4353
16	.5117	.0115	1029	1046	5130	7748	4740	
17			1029				4769	.0958
	.8855	.5036	.7048	.0212	.6243	.6792	.2912	2318
18	, 6514	.4161	-,1304	3012	5269	6979	7691	0153
19 20	.4843 .3179	2003 .0168	2769 2973	1529 1112	8478 2442	8167 7370	2208	.0914
20	.34/9	.0168	29/3	1112	2112	/3/0	1178	.2954
21	. 6939	.5477	.1928	~.1109	5490	7091	5153	.1285
22	.6566	.2348	4005	-,3655	.0271	.1510	1563	7982
23	.7726	.3838	.0718	1482	5544	~.9318	8773	4857
24	.4733	.0982	1047	5509	6163	0333	. 2357	.6388
25	. 6343	.1264	0770	2913	5020	6976	6000	6358
26	. 5945	12721	.0229	2762	4529	0453	3621	63S0
27	,7926	.6245	.5214	.3080	.1390	7024	6334	-,2314
28	.5808	.2813	.3348	. 5366	.7219	.2118	1785	.1409
20	.4174	1192	-,3554	2470	0630	.3718	.6151	1012
30	.8013	.6571	. 5596	.6550	. 6668	.5051	.2830	.0398
31	.4819	.2042	,0990	0625	.3530	.6747	.1589	2385
32	.8240	.6983	.5887	.4578	.0679	,1502	5273	6304
33	.6118	. 1600	2598	5871	6930	6496	.2370	.7217
34	.3599	3901	7004	2503	.0972	.3476	.2104	.1377
35	. 1326	4290	3710	0889	.0975	.0501	4093	2157
36	.1602	1000.	1820	4099	1434	.0279	4352	3010
37	.2388	—,3266	3055	1545	. 2843	.3099	.0568	4240
36	. 5962	.3807	,0039	2860	1000.	.3655	.6985	.5715
35	.7056	.4792	.1504	vaaa	~.0773	-,2769	2707	6200
40	.7307	.4151	.1605	.0755	.2671	.1343	0457	1957
41	.0015	0366	.1109	3358	2600	3133	5200	.4678
42	.5418	,5483	.3090	0074	3748	5073	6030	6047
43	.7378	. 5240	. 2835	0765	1994	4080	6827	6353
44	.7329	.3054	. 1120	0019	3925	8508	8516	4365
45	.4238	.2837	. 2345	1389	2605	← .0031	0387	6898
46	.4462	07×9	3298	4700	5706	3984	.1068	.8165
47	.8593	.8887	.8784	.8208	8589	.7809	.7157	.7478
48	0115	-,1105	.1002	1042	0120	.4323	3705	3478
49	.8189	.2016	.0268	1963	2394	1000.	-,0079	1840
50	.7819	.5547	.5949	.6083	.4/166	.2438	4909	4434
voriigo	. 6027	.2630	.0368	0766	1333	1981	1676	115
(r.)	. 6407	. 3605	.1343	0334	1394	1779	1541	1003
	.8000	, G-J(H)	.5120	.4098	.3277	. 2621	.2097	.1678
-P E(r.)	,1533	. 2795	.3777	.4430	. 4673	. 1100	.3038	.268
vorngo + H.	.7100	.5425	.4345	. 3444	.3310	.2419	.1902	.1520
t drugo + D.	.,,,,,,,	.0123	,1010		.5510		.,,,,,	

Table 1.2. Values of serial cosselation coefficies (r.). Model I-Autoreoffssive model: (i=0.8f.:1+v.

T-35

1 828 2 8556 3 7005 6 8140 6 8140 6 9 7413 9 7413 10 613	. 6105 . 3711 . 4840				•	:	•	•	ė	ī.	-
2 1 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2	. 6105 .3711 .4840		٠				٠	5	:		
. 6725 6 4 1700 6 5 1700 6 6 1700 7 7 7700 8 7410 9 7410 10 6 1810 11 3580	. 6105 . 3711 . 4840 . – . 0671	•	9	•			•	0	11	12	=
2 .0856 4 .3008 6 .8140 6 .004 7 .7792 8 .7419 9 .7419 10 .8181	.3711	.3336	8160.	.0550	1957	.3043	.4035	.6863	.5613	.4427	2113
3 7300 6 3000 6 0000 7 7702 8 7713 9 7713 10 6132	1840	2323	2403	1050	1294	0.283	- 2003	1995	1760	1,160	1
4 . 2005 6 . 8140 0 . 7755 8 . 7755 9 . 7753 10 . 5783	-,0671	1856	0.773	- 0903	0000	100	2555	3843	1.7	A Paris	1000
6		- 3145	- 2000	- 213	0597	1331	1791	6,73	30135	1	1127
6 .0042 7 .7792 8 .7439 9 .7439 10 .6132	. 0095	3750	.0891	1860	4267	1.01	3975	2000	0837	1,0001	10.0481
7 7792 8 7419 9 7419 10 6132 11 3986	.1825	1010	0.00	0.0755	1098	1017	0346	9000	-	2718	1557
8 .7413 9 .7413 10 .6132	6110	0000	- 0264	NIN I	23.18	1382	3031	325.1	1	-	1273
10	8969	4045	3655	96X	0771	1363	0.243	K100	0729	1277	1758
10 .3980	.5888	.3463	2101	.0933	PROG.	100	1 237	- 3303	- 54.35	- 6407	7227
11 3986	. 1642	1.1582	1.344	3334	0.3720	1.094	1,0248	.0905	.3393	.4386	.3703
	0944	0765	-,0380	.0502	.0102	1,337	1195	11131	1802	.0269	- 1758
	6808	.6361	4070	2276	.008	1.11	- 3310	3407	1,4704	5430	- ,5990
13 .6513	DONS.	33.5	2508	-0794	livao.	1610	1010	0701	.0108	1.1057	1773
	. 6437	.5354	. 4363	.3387	1450	010.	-,1517	1.3504	3166	1.214	1,0452
•	.4724	1956	0354	2080	3070	4214	7694	1.43%	- 2135	0010	101
16 .7800	5003	6614	9400	5190	O. WIJ	1876	4226	0219	6197	6830	- 1303
	46644	9.7014	1026	1558	0332	E ONG	200	XIX	02220	3440	- 3133
18 .8156	74911	0477	5850	.5854	4540	4599	3481	35.5	48012	3804	Cipie.
	.6876	0010	. 6207	.4803	3103	.3353	. 4633	. 5167	5529.	.6107	. 6900
	. 1067	1.0558	1760	2049	- 4043	1.3336	1.3748	- 1409	.0014	. I448	7
	. 5964	1010	.1932	0332	0257	0960	14×7		1.108	.1827	.4183
	. 5730	14000	2880	23.38	1045	- OOB:	- 1550	١	1,3790	3417	1,4015
	. 1205	3304	2800	1787	foff37	- 1316	1082		0690	- 0399	0946
24 .8452	6740.	. 5×3G	428	316	3408	Cinia.	2511.5	. 20K7	1960	0000	1070
•	2000	2070	6290	1.0708	1.1071	.0167	.3710	i	.5761	RUNF.	3
sverage .7156	.4852	.3154	1825	81no.	0.10	0407	0326	7200	0156	~ .0503	0571
E(r.) .7413	.5331	.3657	.2312	.1239	.0384	7850	-,0804	1811	1462	1631	1737
90u8*	.6400	.5120	9601	.3277	1206.	.2097	8291.	.1342	. 1073	.0K59	7800
B P E(r.) . (1587	. 1043	.1463	.1784	.2038	12237	.23×4	24.82	.2533	.2535	0617	17.
.versago + 137743	1269	71997	3640	2717	. 20ti7	11077	.215d	2500	6752.	1801.	. 1843

Table 1.3. Values of Berial Correlation coefficients (r_s) Model 11—Autoreoresive model: $\xi_1 = 0.7 \; \xi_{1-1} + 0.6125 \; \xi_{1-2} + v_s$

T = 13

sample no.	۲,	r,		r,	r _e	r.	7,	r.
ī	2	3	4	5	8	7	8	9
1	.2173	5403	6618	0113	.2973	.5173	.5277	2032
<u> </u>	,4106	1726	1026	1543	4620	0×96	.2523	, 2014
3	.3797	2825	0299	. 2069	0828	6××6	7767	2076
i i	.2777	6370	7511	.1182	.7082	.4324	5495	9062
5	.6119	1209	— . 7872	7484	5530	— . 1565	3694	.7710
6	.5006	0325	1741	.1072	.2904	.1425	,1000	1935
7	.3742	— . 1676	0354	.1746	3976	8619	3967	.6236
8	, 5645		1811	.0406	5087	763K	2380	1726
9	. 4254	— . 3621	— . 3×30	0170	.3128	.2467	.0260	-,4759
10	.3497	5272	6907	2720	.2769	.5600	.4540	0933
11	.3881	5841	— . 7969	0727	.7293	.5782	2258	7646
12	.3654	3958	2050	.1189	.1112	2805	5212	2651
13	.3685	5455	75×9	0056	.7765	.4403	3357	7719
14	.3714	2019	to71	.3374	.0074	2355	.1180	. 3391
15	.3917	4709	4213	.8129	.7292	0128	8368	2998
16	. 4585	4040	7005	-,2927	.4495	.5961	.0419	- ,456
17	.0191	-,6626	0757	.3524	.0272	0715	.3010	4773
18	.2188	6567	5203	.3483	.7657	0t13	8126	3704 5126
19	.3741	5010	7411	1987	. 5345	.4477	1245	
20	. 4291	2556	5718	5076	.1013	.4780	.5426	. 2603
21	.4604	1570	2986	3574	2731	.1788	.4725	. 1913
22	.5124	-,4057	-,5657	0551	.2767	, 1984	.0873	1109
23	.1009	5160	.0139	. 1359	3915	3147	. 7306	.2386
24	.4386	3199	4362	. 1082	.2749	.0721	0418	.3906
25	.2303	3441	5038	1058	, 1006	~ .Q018	0847	.1658
26	.3348	5247	-,4838	.2782	.5493	1194	5685	0771
27	.4075	-,1097	1808	1672	4218	6132	4243	.1878
28	.3901	-,2864	-,5694	4165	.0367	.2532	.3079	. 2333
29	.3087	4026	2236	.0008	1426	0568	.0074	3738
30	.4161	3910	-,7017	4212	0841	.3234	.5692	, 60 10
31	.2644	1996	2283	1478	.2279	.3075	2930	2942
32	.3911	5256	8416	3380	.5094	.7736	.1112	— , Bikid
33	.5096	1010	7127	7886	2045	. 1993	.5078	.2817
34	.5420	2937	4277	0923	1117	1299	,0680	. 3353
35	.3217	-,1926	.3750	,2844	.1329	.3190	.3084	5999
36	.5298	1262	5218	6331	5686	3364	.0640	.2127
37	.4411	3208	—.6707	2104	, 3367	.4947	.4118	071×
38	.3058	-,4731	7216	~.1980	.4990	.4822	.0799	3298
39	.4311	2312	→,6014	2794	.0398	.0249	,1796	. 5254
40	.1063	— .8351	- , 3555	. 6363	.4893	4490	— .66N0	.2339
41	.4162	1419	64 40	~.7117	-,1384	.4404	.8910	, 3666
42	.5949	0252	6319	7136	5923	1406	. 2755	. 5954
43	.1701	7562	4201	.3326	. 5532	1245	6956	0589
44	.4015	1518	.1249	.1490	-,3040	4976	2099	37×0
45	.0159	2330	1290	2624	3657	. 4660	.5613	.0291
46	.2081	3156	3396	.0416	.0989	.0489	1703	÷.4310
47	.3263	4114	6386	4436	.3497	.4709	.1320	-,1816
48	.2564	— . 5768	5430]	.0956	.4897	.1193	2787	0414
49	.4372	4386	8030	— .3701	.2045	.7040	.5840	.1237
50	.2925	7049	5543	.1086	. 6030	.6655	1469	8427
everage	.3617	3614	4387	0693	.1274	.1046	.0151	0750
£(r,)	.4000	3678	8186	1921	.1732	.2110	.0104	1056
ρ,	.4341	3086	4809	1483	.1014	.2248	.0102	1096
$B_* = P_* - E(r_*)$.0251	.0592	.0677	.0438	.0182	.0138	0002	0010
average + B.	.3868	3052	3710	0455	. 1456	.1184	.0149	0760

Table 1.4. VALUES OF BERIAL CORDELSTON CORPUCENTS (r.). Moure 11.—Autorresembery Moure, (i.-0.1f., - 0.0123 f., + v.

wmple no.	-				٠	.*	ē		٠	2	3	13
-	64		•	10	20		*	a	10	=	2	= =
-	.3162	3728	3430	.1526	.2830	- 1633	F6-06	00,200			1	!
64	4400	1.3459	1.4439	0857	2691	25.5	1070	200	Eco .	1.0555	1.207	1.2565
•	2460	0110.	1.2230	- 1354	1	187.	1000	1	-	.6043	6516	. 135
7	7008.	4944	1.0505	0836	1662	1731		3	12183	-,0895	100	0010
49	. 6972	.1738	1.1661	2013	1000	0566	125	2807	1 1 1 1 1 1	1331	52.80.	9280
•	8002	0507		4000	*							199.1
•		1		1.0832	.0240	1070.	2761	1,2913	1497	2477	2110	1543
- :	9779	1	SCEO.	.3707	.0393	1776	1530	.1373	0.03	0.1178	-	
	100	10.1	1888	176	1.2480	1.1434	0040	1052	0.354	250	-	
	3792	1.1838	1.3850	2018	1280	.1693	. 1337	1,1064	- 6:23	4103	4777	1000
2	.3394	- , 5387	ets	0146	4002	3656	0571	2182	0213	0000	- 1810	1205
=	4911	1001	1,11	0110		į	-					
::	920		1	20.00	BCCC.	. 2773	- 080°	1.2038	.0513	. 193\$	0530	196.1
::	100					. 2514	1.3511	6405.1	1297	4049	3	1460
?:	00:1				* 619		1.2246	R66+	1021.	.3710	.3148	1.1564
==			1	1.5	700	4587	. 1951	1.2167	1.45	1000	13:15	202a
2		133	1.0038	.1308	.0083	.0343	5100	.0585	7.0760	1,1080	1.063	.0267
9	.2356	0807	1165	149	9	2000	4010	911	9	-	, 0	
11	. 2920	- 4377	100	6756	200	0000		1.0.100	5	5000		2
18	. 1950	2347	1.2784	0153	1831	1083	7:00:1	9	200	1 4.94	200	000
19	3×30	3948	- 8300	2333	216	1800	10565				200	1
50	4702	1.1874	5:34	1909	1000	1721	.0311	- 1184	1,1086	- 0×03	i	3381
- 5	.3787	3618	1.5041	2816	.1422	:3952	.1793	2613	3350	0107	.2552	.0795
71	.6275	2578	1585.	2067	3509	4440	.0523	1330	2318	2645	6:17	- 0427
ti	.3935	3567	4837	1.3420	1.2679	0020	4980	31111	1	1557	1219	1200
71	.4146	1.3021	1.4664	1.000	71	10.08	1.4.155	1	41.54	3580	1	4900
52	.2789	1.3813	3605	.0317	1976.	.2748	1910	1.41	1070	.3300	7.75	1.2015
avoraga	1202.	2750	3811	0877	.1565	.1079	Otrio -	1720	90800-	8090	.0545	-,0492
E(r.)	.4241	3320	1809.	1683	.1784	.2127	.0245	-,1281	1191	0185	2740.	.0324
	.4341	3080	4819	1483	1911	3248	.0402	-,1096	-, 1013	1.0038	.0394	.0439
B P E(r.) ,0100	r.) ,0100	.0234	.0265	.0103	.0130	1210.	.0157	.0185	.0178	.0147	.0122	.0115
avorage + B.	,4021	2516	3540	8790	.1605	.1200	1.031€	1.1541	0728	.0756	1000.	0377

Table 2.1. Frequency distribution of x. tor texting tree along cance of difference between observed and theoretical strial correlations

Autorograssive Model : { -0.7 { -1+0.0125 { -1 = 7.

Autorogramive Model : $\xi_1 = .8\xi_{1-3} = \eta_1$

iongth of model sample = 35 Iongth of model sample	longth of model sample	longth of model sample	longth of model sample	e oldum lopour Jo unblo	lol sumplo =		\$1.	length sample	longth of model	longth semp	longth of model
x'forr, x'forr, x'forR, Ex'forR, x'forr, x'forr, x'forR,	Ex! for R. x' for r, x' for r,	Ex! for R. x' for r, x' for r,	x, for r,		5 . 7 .		TX for R. X for R.	X, for R,	, X	x' for R.	KX.
d.f.=1 d.f.=1 d.f.=10 d.f.=1 d.f.=1 d.	d.f 10 d.f 1 d.f 1	d.f 10 d.f 1 d.f 1	d. f. = 1		· i	d. f.	d. f. = 0	d. f. = 1	d. f. == 12	d. f.= 12 d. f.= 1	d. f. == 8
3 4 5 6 7 8	5 6 7 8	6 7 8	7 8	8			9	=	13	2	7
1 1 2	1 2 1	1 2	51 	ı		73	,	-	ı	e	1
1	1	1	1	1		-	1	•	ı	64	ı
2 9 I		-	-	-		ø	I	10	-	۲-	ļ
1 1 10 2 4 -	10 2 4 -	7 *	1	1		7	64	4	ı	13	1
2 4 19 - 10 6	19 - 10 6	10 6	10 6	ø		18	•	25	-	19	I
6 3 21 2 5 4	21 2 5 4	2 6	9	•		18	-	18	7	92	ı
4 6 35 4 8 6	35 4 8	8 8	8	80		Ş	-	Ŧ	6	37	eı
5 3 48 3 7 10	48 3 7 10	3 7 10	7 10	91		70		48	i	\$	*
3 26 2 4 8	20 2 4 8	° ° ° ° ° ° ° ° ° ° ° ° ° ° ° ° ° ° °	8	80		R		20 11	I	33	I
2 3 33 - 3 8	33 - 3 8		8	00		\$	7	35	61	4.1	•
- 2 10 1 1 2	10 1 1 2	1 1 2	61	eı		11		16	64	33	61
9 1	6 3 1 3	3 1 3		e.		61	61	22	ı	23	4
1 5 2 2	5 2 2 1	2 2 1	2 1	-		80	-	G	61	21	-
1 1 8 1 1	11 6 11	1 1	- 1			54	11	33	13	90	35
25 25 250 25 60 60	25 60	99		92		300	20	300	22	400	3

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Table 2.2. Values of χ^2 given by Quenouille's test for goodness of fit of a phiori enown models to aterage correlogram Model: $\{\iota_1=0.8\,\xi_{i+1}=\tau_i$

_	longth	35	longth	15
•	R.	x'	R_{\bullet}	יא
ı	2	3	4	5
	w	ith correction	n for bise	
1	0257	1.58	0840	13.72*
2	.0099	.62	.0247	3.00
3	.0008	.00	.0184	1.56
4	acoo. ~	.57	.0258	2.83
5	,0028	.04	0580	12,984
G	.0409	9.34*	.0220	1.82
7	.0316	5.37*	0065	.13
8	.0322	5.39*		
9	0250	3.13		
10	0217	2.27		
11	.0187	1.62		
otal		29.91		36.10

ı	0844	16.82*	2373	100.40
2	→ .0029	-05	0039	.07
3	0110	.88	.0008	.00
4	0217	2.83	.0256	2.78
5	0000	.47	0338	4.414
G	.0300	5.02*	.0640	14,24*
7	.0216	2.53	.0267	2.03
8	.0234	2.87		
9	0321	5.18*		
10	0271	3.53		
11	.0134	.83		
tolal		41.01		133.02

[·] significant at 5% level

Table 2.3. Values of X³ given by Quenouille's test for goodness of sit of a trior; enown models to average correlogram

Model: \$1-0.7 \$1., +0.6125 \$1.2 = 4.

	Montel: (-0.7 (+1	J.6120 E V	•
4-2	longt)	35	lengtl	15
1-2	R,	X*	R.	x'
1	2	3	4	5
	٧	rith correctio	n for bian	
1	0193	1.24	.0082	.10
2	.0275	2.43	0048	.05
1	.0229	1.63	0218	1,11
4	0230	1.73	.0410	3.59
6	0034	.03	0017	.00
В	0474	6.34*	0355	2.21
7	.0158	.68		
8	.0021	.01		
9	0509	6.56*		
10	.0030	.02	••	••
total		20.66		7.08
	•	rithout corre	tion for bias	
	₹.0340	3.83*	0300	2.45
2	.0132	. 34	-,0350	3.10
3	.0088	. 23	051B	6.25
4	0380	4.35*	.0126	.34
٥	0173	.88	0009	.00
6	0610	10.40*	→ .0593	5.15°
7	.0025	.01		
8	0105	.29		
9	0831	10.08*		
10	0083	.17	••	••
total		30.67		18.29

^{*} significant at 5% level

Table 4. FREQUENCY DISTRIBUTION OF X' AS GIVEN BY QUENOUILLE'S TEST FOR SERIAL CORRELATION MODEL FITTED BY DIFFERENT METHODS; LENOTH OF THE SAMPLE = 35,

Model: {++a, {++} + a, {+-} = *. expected percentage frequency fitted from r, and r, fitted by minimising $\sum_{n=1}^{12} (r_n + a_1 r_{n-1} + a_1 r_{n-1})^2$ P (x1) x* for R. cumulative X1 χ' for R, cumulative d.f.=1 for R. d.f.=1 χ' for R, d.f.=10 6 2 3 4 δ .98 .95 .90 .80 .70 .50 .20 .10 .05 .02 3 1 1 3 1 2 1 3 5 10 10 20 10 10 5 3 1 1 13 16 27 22 38 51 20 21 13 15 2 1 3 4 1134231 30 24 45 44 10 27 8 5 3 14 2 .01 ıï total 100 250 25 250 25

Table 3. Values of a₁, and b₂ obtained by dispersit methods of pitting time string from model: $\xi_1 - 0.7 \, \xi_{1.1} + 0.0126 \, \xi_{1.2} = \eta_1$

*ample		From	from r, and ra		minimi	minimising \(\mathbb{Z}(r_a + \alpha_1 r_{a+1} - a_a r_{a+1})^4\)	0 1 7.1 - 0	۱۰ ۲۰-۱)،		minimi	minimizing 2.R.	
	ű	% diff.	4	% diff	ē	% diff	i	% diff.	i	% diff.	£	% diff.
-	21		7		٩	7	100	•	10	=	22	12
mortol	-0.7		0.0125		-0.7		0.0123		1.0-		0.6125	
-	- 4823	31.1	. 5233	1.3	4589	34.4	.6120	0.0				
ŧı	1584	20	GING.	15.1	7846	-	.8104	32.3				
n	1,1683	1:1	4447	27.4	7649	8.3	3	51.2	1.3729	16.7	781	29.1
+	- 6300	9.0	.7310	10.3	7637	1,0	8433	37.7				
••	7,7400	0.0	184	53.4	0.500	37.6	1530	30.0				
Ð	7227	3.5	4239	30.5	1141	6.9	.4038	34.1				
7	-, 5995	1.4	3474	37.6	4437	30.6	.3203	47.7				
30	×340	19.3	Cint.	17.3	8299	18.6	.4517	20.3	6143	***	3101	48.4
9	5713	18.4	Pinty.	11.3	7166	?;	08.29	11.5				
10	5902	15.7	.7390	20.1	6526	8.8	7300	19.3				
=	7510	7.3	.5392	13.6	7100	1.4	.5934	3.1				
2	3550	40.4	DIXT.	30.4	-, 4573	34.7	1191	0.02				
2	1,6790	£.	.x172	33.4	- 6000	÷	.8447	33.15	-1.0603	81.8	3.	9.9
Ξ	1.7554	7.0	. 1575	23.7	1.7931	13.3	.7710	90.07				
13	1.2500	63.3	. 1945	63.5	2724	1.1	4308	6-:-3				
16	1112	61.3	Days.	75.4	1.3289	53.0	.2047	60.6				
11	1 460	34.5	. 5724	6.5	×111×	34.10	.8130	18.2				
×	1.2515	- 19	Dt M.S.	53.4	3330	52.4	4138	7:1	- 2013	02.7	.4181	31.7
=	1.6260	10.6	. 0343	3.6	- 4713	7	64143	20.20				
OÇ.	7108	;	. 5244	1.4	8430	20.4	644	4.8				
53	1.6021	14.0	58818	3.7	7034	0.5	05.85	11.8				
?;	0103	31.3	.7427	21.3	1003	13.0	.7750	0.83				
53	6317	20.00	. 43133	1	7731	10.	SAKE.	0.4	670	**	. 5513	20.0
₹,	1.6813	 9	K719.	10	- 6103	2.	7293	9.3				
£1	4177	40.3	8104	18.7	1555	25.4	, UN17	11.3				
averave.	- 6017	13.6	5304	2	- 6113	7	5841	-				

GOODNESS OF FIT OF A MOVING AVERAGE WITH THERE CONSTANTS FITTED TO AUTOMORPHINE SKILLES OF LENGTH 35 TABLE 5. VALUES OF X!

							•	CONTERNI	AUTOGRESSIVE BEILIES OF LESOTE 35	5	and the	2						
	sample no. I	no. F			entaple tto. 2	10. 2		sample no. 14	no. 14		Memple no. 15	0.13		Munple no. 18	no. 18		memple no. 17	10.17
lag no.		, x	1,4	ag e	, x	xx.	Ž 5	ž	ž	<u>1</u>	*	ž,	¥ è	×	ž,	1 5 E		×
-		74	-	*	•	æ	-	30		2	 = 	13	=	=	15	2	=	*
•	4 .4213	213	.4213	•	4 .1117	1111	*	.2730	.2730	-	4016	1194.	-	.5397	1659.	*	\$050.	1020
*	5 1.0272	272	1.4485	10	1.7237	6 1.7237 1.8354	4	5 2.3701 2.6521	2.0521	10	.0817	.6433	0	.1273	.6870	10	.691	.1116
۰	**	\$.6058 ♠	4.0543	9	6760.	1.8727	9	.2068	2.8589	ъ	.0269	.6720	9	1290	.7491	9	.0007	.1123
•	ij	.3083 4	4.3026	7	7 2.8053	4.6780	7	. 6323	3.4914	4	.1058	.6700	-	1782	1.0273	1	3	.6134
-	•	\$100.	4.3644	90	8 1.5345	6.2125	40	.001	3.4925	90	.0407	.7267	90	.0014	1.0287	•	.3974	1.0108
•	۰.	6770.	4.4417	œ	.6558	0.8683	e	.3083	3.8908	•	2779	1.0036	۰	.0210	1.0497		.1218	1.1326
2	۰.	.0750	4.6176	01	1.7020	8.5709	2	.1070	3.9678	2	.2940	1.2076	01	0000.	1.0797	2	.0704	1.8030
=	: :	2.1806	6.6082	=		1291 8.7000	Ξ	.2080	4.2358	Ξ	0000	1.3055	=	.0344	1.1141	=	6130	2.0220
12	Ψ,	1586	12 .0586 7.3568	-	1,3454	12 1, 1454 10 0454	2	13 Dins 4 2766	4 9766	:	5	9000	•	0,11	EZILE 1995. 91 1799 1 OCIT #1 880F 1 9050 91	13	4947	3.1173

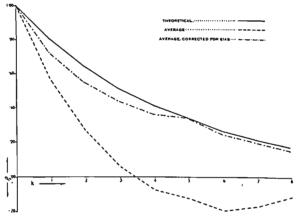


Fig. (1.1). Autoregressive Model I. T=15

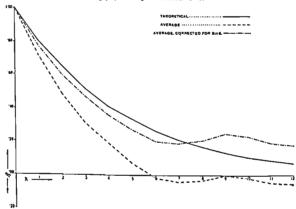


Fig. (1.2). Autoregressive Model I. T=35

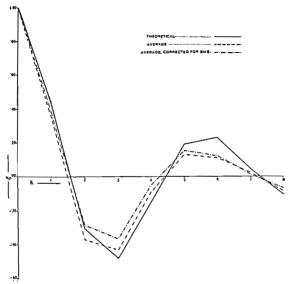


Fig. (1.3). Autoregressive Model II: T=15

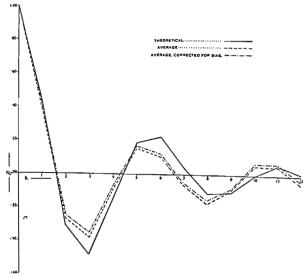


Fig. (1.4). Autoregressive Model II : T=35