# BOUNDS ON MOMENTS OF CERTAIN RANDOM VARIABLES 

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1. Summary. Let $\left\{X_{n}, n \geqq 1\right\}$ be a sequence of random variables and let $S_{n}=\sum_{i=1}^{n} X_{i}$. Under the condition that $\left\{S_{n}\right\}$ forms a martingale sequence, it was shown in [2] that, for $\nu \geqq 2$,

$$
\left.E\left(\mid S_{n}\right)^{\nu}\right) \leqq C_{\nu} n^{(\nu / 2)-1} \sum_{i=1}^{n} E\left|X_{i}\right|^{\nu},
$$

where

$$
\begin{equation*}
C_{\nu}=\left[8(\nu-1) \max \left(1,2^{\nu-3}\right)\right]^{\nu} . \tag{1}
\end{equation*}
$$

The purpose of this paper is to show that the constant $C_{\nu}$ can be replaced by a much smaller constant in the following two cases: (i) $\nu$ is an even integer and the martingale dependence condition is replaced by one which is more explicit in terms of moments (Theorem 1); (ii) the $X_{n}$ 's are independent with zero means (Theorem 2). For case (i) we give for $E\left(\left|S_{n}\right|^{\nu}\right)$ a bound which is a polynomial in $n$. This last bound does not appear to be too exhorbitant because, as shown by an example, it is not valid for all martingales $\left\{S_{n}\right\}$.
2. The results. We first prove the following

Theorem 1. Suppose that for every integer $p \geqq 1$ and for every choice of positive integers $i_{1}, \cdots, i_{p}, k_{1}, \cdots, k_{p}$, the condition $\min \left(k_{1}, \cdots, k_{p}\right)=1 \Rightarrow$ $E\left(X_{i_{1}}^{k_{1}} \cdots X_{i_{p}}^{k_{p}}\right)$, if it exists, equals zero. Then, for $m=1,2, \cdots$,

$$
E\left(\left|S_{n}\right|^{2 m}\right) \leqq D_{2 m} n^{m-1} \sum_{i=1}^{n} E\left|X_{i}\right|^{2 m},
$$

where

$$
\begin{equation*}
D_{2 m}=\sum_{p=1}^{m} p^{2 m-1} /(p-1)!. \tag{2}
\end{equation*}
$$

Proof. To make the writing simpler we write $\gamma_{\nu, n}=E\left|X_{n}\right|^{\nu}$ and $\beta_{\nu, n}=$ $\sum_{i=1}^{n} \gamma_{\nu, i} / n$. Keep $n$ and $m$ fixed. The result holds if $\beta_{2 m, n}=\infty$.

Suppose therefore that $\beta_{2 m, n}<\infty$. For $1 \leqq p \leqq 2 m$, let $A_{p}$ denote the set of all $p$-tuples $\mathbf{k}=\left(k_{1}, \cdots, k_{p}\right)$ such that the $k$ 's are positive integers satisfying $\left(k_{1}+\cdots+k_{p}\right)=2 m$. Let

$$
T\left(i_{1}, \cdots, i_{p}\right)=\sum(2 m)!/\left(k_{1}!\cdots k_{p}!\right) E\left(X_{i_{1}}^{k_{1}} \cdots X_{i_{p}}^{k_{p}}\right)
$$

where the summation is over $\mathbf{k} \varepsilon A_{p}$. Then

$$
\begin{equation*}
E\left(\left|S_{n}\right|^{2 m}\right)=E\left({S_{n}^{2 m}}^{2 m}\right)=\sum_{p=1}^{2 m} \sum^{\prime} T\left(i_{1}, \cdots, i_{p}\right), \tag{3}
\end{equation*}
$$

where $\sum^{\prime}$ denotes summation over the region $1 \leqq i_{1}<\cdots<i_{p} \leqq n$. If $p>m$ and $\mathbf{k} \varepsilon A_{p}$ then $\min \left(k_{1}, \cdots, k_{p}\right)=1$. Thus $p>m \Rightarrow>T\left(i_{1}, \cdots, i_{p}\right)=0$.

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Moreover by Hölder's inequality

$$
\left|E\left(X_{i_{1}}^{k_{1}} \cdots X_{i_{p}}^{k_{p}}\right)\right| \leqq \gamma_{2 m, i_{1}}^{k_{1} / 2 m} \cdots \gamma_{2 m, i_{p}}^{k_{p} / 2 m} .
$$

Therefore

$$
\begin{aligned}
\left|T\left(i_{1}, \cdots, i_{p}\right)\right| & \leqq\left(\gamma_{2 m, i_{1}}^{1 / 2 m}+\cdots+\gamma_{2 m, i_{p}}^{1 / 2 m}\right)^{2 m} \\
& \leqq p^{2 m-1}\left(\gamma_{2 m, i_{1}}+\cdots+\gamma_{2 m, i_{p}}\right)
\end{aligned}
$$

Thus, from (3),

$$
\begin{aligned}
E\left(\left|S_{n}\right|^{2 m}\right) & \leqq \sum_{p=1}^{m} p^{2 m-1} \sum^{\prime}\left(\gamma_{2 m, i_{1}}+\cdots+\gamma_{2 m, i_{p}}\right) \\
& =\sum_{p=1}^{m} p^{2 m-1}\binom{n-1}{p-1} \sum_{j=1}^{n} \gamma_{2 m, j} \\
& \leqq \sum_{p=1}^{m} p^{2 m-1} n^{p-1} /(p-1)!\cdot n \beta_{2 m, n} \\
& \leqq n^{m} \beta_{2 m, n} D_{2 m} .
\end{aligned}
$$

This completes the proof of the theorem.
Remark 1. Berman [1] calls a sequence $\left\{X_{n}\right\}$ of random variables sign-invariant if, for every $n$ and for every choice of $\epsilon_{1}, \cdots, \epsilon_{n}$ each equal to +1 or -1 , the joint distribution of $\epsilon_{1} X_{1}, \cdots, \epsilon_{r} X_{n}$ is the same as that of $X_{1}, \cdots, X_{n}$. It is easy to see that the moment condition of the above theorem is satisfied if $\left\{X_{n}\right\}$ is sign-invariant. It also holds if the $X_{n}$ 's are independent with zero means.

Theorem 2. Suppose that the $X_{n}$ 's are independent random variables with zero means. Then, for $\nu \geqq 2$,

$$
E\left(\left|S_{n}\right|^{\nu}\right) \leqq F_{\nu} n^{\nu / 2-1} \sum_{i=1}^{n} E\left|X_{\imath}\right|^{\nu},
$$

where

$$
F_{\nu}=\frac{1}{2} \nu(\nu-1) \max \left(1,2^{\nu-3}\right)\left[1+2 \nu^{-1} D_{2 m}^{(\nu-2) / 2 m}\right]
$$

where the integer $m$ satisfies $2 m \leqq \nu<2 m+2$ and the constant $D_{2 m}$ is given by $(2)$.
Proof. We will use the notation introduced at the beginning of the proof of Theorem 1. Keep $\nu$ and $n$ fixed. Again the result holds if $\beta_{\nu, n}=\infty$. Suppose therefore that $\beta_{\nu, n}<\infty$.

Let $\Delta_{n}(\nu)=E\left(\left|S_{n}\right|^{\nu}-\left|S_{n-i}\right|^{\nu}\right)$. Then from the relation (3.5) of [2], we get

$$
\begin{equation*}
\Delta_{n}(\nu) \leqq \frac{1}{2} \nu \delta_{\nu}\left[\gamma_{2, n} E\left(\left|S_{n-1}\right|^{\nu-2}\right)+\gamma_{\nu, n}\right], \tag{4}
\end{equation*}
$$

where $\delta_{\nu}=(\nu-1) \max \left(1,2^{\nu-3}\right)$. Now $\nu-2<2 m$. Therefore, using Theorem 1 , we get

$$
\begin{align*}
E\left(\left|S_{n-1}\right|^{\nu-2}\right) & \leqq\left[E\left(\left|S_{n-1}\right|^{2 m}\right)\right]^{(\nu-2) / 2 m}  \tag{5}\\
& \leqq D_{2 m}^{(\nu-2) / 2 m}(n-1)^{(\nu-2) / 2} \beta_{2 m, n-1}^{(\nu-2) / 2 m} .
\end{align*}
$$

Using the inequality $\beta_{2 m, n-1} \leqq \beta_{\nu, n-1}^{2 m / \nu}$ in (5) and the inequality $\gamma_{2, n} \leqq \gamma_{\nu, n}^{2 / \nu}$ in (4), we obtain

$$
\begin{equation*}
\Delta_{n}(\nu)=\frac{1}{2} \nu \delta_{\nu}\left[D_{2 m}^{(\nu-2) / 2 m}(n-1)^{(\nu-2) / 2} \beta_{\nu, n-1}^{(\nu-2) / \nu} \gamma_{\nu, n}^{2 / \nu}+\gamma_{\nu, n}\right] . \tag{6}
\end{equation*}
$$

From the corollary to Lemma 2 of [2], we have

$$
\begin{equation*}
\sum_{j=1}^{n}(j-1)^{(\nu-2) / 2} \beta_{\nu, j-1}^{(\nu-2) / \nu} \gamma_{\nu, j}^{2 / \nu} \leqq 2 \nu^{-1} n^{\nu / 2} \beta_{\nu, n} \tag{7}
\end{equation*}
$$

Finally, inequalities (6) and (7) yield

$$
\begin{aligned}
E\left(\left|S_{n}\right|^{\nu}\right) & =\sum_{j=1}^{n} \Delta_{j}(\nu) \\
& \leqq \frac{1}{2} \nu \delta_{\nu}\left[D_{2 m}^{(\nu-2) / 2 m} 2 \nu^{-1} n^{\nu / 2} \beta_{\nu, n}+n \beta_{\nu, n}\right] \\
& \leqq n^{\nu / 2} \beta_{\nu, n} F_{\nu} .
\end{aligned}
$$

This completes the proof of the theorem.
Remark 2. Suppose that the $X_{n}$ 's are independent with zero means. Then $\left\{S_{n}\right\}$ is a martingale and the Theorem of [2] is applicable. However, Theorem 2 above also applies and gives a better bound. If, moreover, $\nu$ is an even integer, then Theorem 1 gives a still better bound.

Remark 3. Let $\left\{X_{n}, n \geqq 1\right\}$ be an exchangeable process with $E\left(X_{1} X_{2}\right)=0$. Then the proof of Theorem 2 breaks down. However, the conclusion is valid because of the de Finetti theorem (see Section 4 of [2]).

Remark 4. Let $\beta_{\nu, n}^{\prime}=\max \left\{E\left|X_{i}\right|^{\nu}, 1 \leqq i \leqq n\right\}$. If the moment condition of Theorem 1 holds then minor modifications of the proof of that theorem show that

$$
\begin{equation*}
E\left(\left|S_{n}\right|^{2 m}\right) \leqq \beta_{2 m, n}^{\prime} \sum_{p=1}^{n}\binom{n}{p} \sum^{\prime \prime}(2 m)!/\left(k_{1}!\cdots k_{p}!\right) \tag{8}
\end{equation*}
$$

where $\sum^{\prime \prime}$ denotes summation over the region $k_{i} \geqq 2$ and $k_{1}+\cdots+k_{p}=2 m$. Thus

$$
\begin{aligned}
& E\left(\left|S_{n}\right|^{2 m}\right) \leqq \beta_{2 m, n}^{\prime}\left[(2 m)!/\left(2^{m} \cdot m!\right) n^{m}\right. \\
& \left.\quad+(2 m)!(m-5) /\left(9 \cdot 2^{m} \cdot(m-2)!\right) n^{m-1}+o\left(n^{m-1}\right)\right]
\end{aligned}
$$

Note that the leading term has coefficient $(2 m)!/\left(2^{m} \cdot m!\right)$, which is natural in view of the central limit theorem.

Remark 5. Suppose the $X_{n}$ 's satisfy the moment condition of Theorem 1 and in addition are identically distributed. Then the bound (8) is better than that given by Theorem 1.

Remark 6. The bound (8) does not appear to be exorbitant in that it is not valid for all martingales, as seen from the following example. Let the basic probability space be $\{1, \cdots, 6\}$ with the points $1,2,5$ and 6 getting the mass $\frac{1}{8}$ each and the points 3 and 4 getting the mass $\frac{1}{4}$ each We define three random variables $X_{1}, X_{2}$ and $X_{3}$ on the space with values given in the following table.

|  | Point |  |  |  |  |  |
| :---: | :---: | ---: | :---: | :---: | :---: | :---: |
| Random <br> Variable | 1 | 2 | 3 | 4 | 5 | 6 |
| $X_{1}$ | 1 | 1 | 1 | -1 | -1 | -1 |
| $X_{2}$ | 1 | 1 | -1 | 1 | -1 | -1 |
| $X_{3}$ | $2^{\frac{1}{4}}$ | $-2^{\frac{1}{4}}$ | 0 | 0 | $2^{\frac{1}{4}}$ | $-2^{\frac{1}{2}}$ |

The sequence $S_{1}, S_{2}, S_{3}$ of partial sums forms a martingale. Further $E S_{3}{ }^{4}=$ $9+12 \cdot 2^{1 / 2}$ which exceeds the bound 21 , given by (8).

## REFERENCES

[1] Berman, S. M. (1965). Sign-invariant random variables and stochastic processes with sign-invariant increments. Trans. Amer. Math. Soc. 119 216-243.
[2] Dharmadhikari, S. W., Fabian, V. and Jogdeo, K. (1968). Bounds on the moments of martingales. Ann. Math. Statist. 39 1719-1723.

