E-Optimal Minimally Connected Block Designs Under Mixed Effects Model

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Summary: Considering a mixed effects model in a minimally connected block design set-up, we obtain designs which are E-optimal, uniformly in the ratio of the variance components, for inference on varietal contrasts which constitute the fixed effects in the model.

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1 Introduction and Preliminaries

A connected block design is said to be *saturated* or *minimally connected* if the error degree of freedom is zero under a fixed effects model. This happens when the usual design parameters $v, b, k (\geq 2)$ satisfy the relation

$$v = b(k-1) + 1 (1.1)$$

The study of optimal minimally connected block designs is fairly recent and the following results are known under a fixed effects additive model without interaction:

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- (a) all minimally connected designs are D-equivalent;
- (b) any design having one treatment common to all the blocks is A- and E-optimal; such optimal designs are isomorphic to one another and they are uniquely A- and E-optimal.

Throughout this paper, two designs are said to be isomorphic to each other if any one of them can be obtained from the other by renaming the blocks and/or treatments. We refer to Mukerjee, Chatterjee and Sen (1986), Krafft (1988), Mandal, Shah and Sinha (1991) and Dey and Bapat (1991) for the above-mentioned and various other related results. Two related references are Mukerjee and Sinha (1990) and Birkes and Dodge (1991).

In this paper, we initiate a study of optimal minimally connected block designs under a mixed effects model where the block effects are random with variance σ_b^2 and the treatment effects are fixed. For fixed v = b(k-1)+1, $b, k (\geq 2)$, let $\mathcal{D} = \mathcal{D}(v, b, k)$ be the class of all minimally connected designs in the sense of (1.1) and, as usual, let σ_ϵ^2 denote the error variance. Define $\theta = \sigma_b^2/\sigma_\epsilon^2$. Then for each design $d \in \mathcal{D}$, the information matrix for the treatment effects is (see Rao (1947), Shah and Sinha (1989, p. 86))

$$C_d = C_{1d} + (1 + k\theta)^{-1} C_{2d} , (1.2)$$

where

$$C_{1d} = R_d - k^{-1} N_d N'_d$$
, $C_{2d} = k^{-1} N_d N'_d - (bk)^{-1} r_d r'_d$, (1.3)

 N_d is the incidence matrix of d, $R_d = \text{diag}(r_{1d}, \ldots, r_{vd})$, $r_d = (r_{1d}, \ldots, r_{vd})'$, r_{1d}, \ldots, r_{vd} being the replication numbers for the v treatments in d. Note that $\theta = 0$ corresponds to no differential block effects while $\theta = \infty$ corresponds to the fixed effects model. Our study reveals that the design outlined in (b) above continue to be E-optimal uniformly in θ .

2 E-Optimal Designs and their Uniqueness

Lemma 2.1: Each $d \in \mathcal{D}$ contains at least two blocks such that each of these blocks contains k-1 treatments which are replicated exactly once in d.

Proof: Let $T_d = \{i: 1 \le i \le v, r_{id} = 1\}$. If m_{1d} is the cardinality of T_d then one must have $b \nmid k \ge m_{1d} + 2(v - m_{1d})$, so that by (1.1),

$$m_{1d} \ge b(k-2) + 2$$
 (2.1)

If m_{2d} is the number of blocks in d containing k-1 members of T_d then considering the occurrence of the members of T_d in d,

$$m_{1d} \le m_{2d}(k-1) + (b-m_{2d})(k-2) = b(k-2) + m_{2d}$$
.

Hence by (2.1), $m_{2d} \ge 2$, completing the proof.

In consideration of Lemma 2.1, for each $d \in \mathcal{D}$, rearranging the treatments and blocks, the incidence matrix N_d can be expressed as

$$N_d = \begin{pmatrix} 1_{k-1} & 0 & 0 \dots 0 \\ 0 & 1_{k-1} & 0 \dots 0 \\ \hline & & \\ \hline & & \\ N_{1d} \end{pmatrix} , \qquad (2.2)$$

say, where for positive integral q, 1_q is a $q \times 1$ vector with all elements unity. Defining the $v \times 1$ vector $x = (1'_{k-1}, -1'_{k-1}, 0')'$ it follows, after some algebra, from (1.2), (1.3), (2.2) that for each $d \in \mathcal{D}$,

$$x'C_dx = (x'x)\{k^{-1} + (1+k\theta)^{-1}(1-k^{-1})\}$$
.

Therefore, for each $d \in \mathcal{D}$, the minimum non-zero eigenvalue of C_d satisfies

$$\lambda_{\min}(C_d) \le k^{-1} + (1 + k\theta)^{-1} (1 - k^{-1}) . \tag{2.3}$$

Consider now the design $d^*(\in \mathcal{D})$ with blocks $\{1, 2, ..., k\}$, (1, k+1, ..., 2k-1), ..., $\{1, (k-1)(b-1)+2, ..., (k-1)b+1\}$. By explicit computation using (1.1)-(1.3), it can be seen that the non-zero eigenvalues of C_{d^*} are

$$\lambda_1^* = 1$$
 , $\lambda_2^* = k^{-1} + (1 + k\theta)^{-1} (1 - k^{-1})$, $\lambda_3^* = k^{-1} v$, (2.4)

with respective multiplicities b(k-2), b-1 and 1. Since $\theta \ge 0$, one obtains

$$\lambda_3^* > \lambda_1^* \ge \lambda_2^* \tag{2.5}$$

Hence $\lambda_{\min}(C_{d^*}) = \lambda_2^*$, and comparing this with (2.3), the following holds:

Theorem 2.1: For each θ ($0 \le \theta \le \infty$), designs isomorphic to d^* are E-optimal in \mathcal{D} .

In the rest of this section, we study the uniqueness of the *E*-optimal designs mentioned in Theorem 2.1. The following lemmas will be helpful.

Lemma 2.2: Suppose $d \in \mathcal{D}$ is not isomorphic to d^* . Then d contains two blocks such that

- (i) each of these two blocks contains k-1 treatments which are replicated exactly once in d, and
- (ii) the remaining treatments in these two blocks are distinct.

Proof: Let T_d be as in the proof of Lemma 2.1, B_d be the set of blocks in d which contain exactly k-1 members of T_d , and $a=a_d$ be the cardinality of B_d . By Lemma 2.1, $2 \le a \le b$. If possible, suppose the conclusion of the lemma is false. Then there is some treatment common to each block of B_d . In that case, without loss of generality, let the a blocks in B_d be $\{1, 2, \ldots, k\}$, $\{1, k+1, \ldots, 2k-1\}$, ..., $\{1, (k-1)(a-1)+2, \ldots, (k-1)a+1\}$.

Since d is not isomorphic to d^* , one gets a < b, i.e., the set, \vec{B}_d , of blocks of d, which are not members of B_d , is non-empty. Define

$$\bar{T}_d = \{i: (k-1)a + 2 \le i \le (k-1)b + 1, r_{id} = 1\}, \quad \bar{m}_{1d} = \text{ cardinality of } \bar{T}_d$$

 \bar{m}_{2d} = number of blocks in \bar{B}_d containing exactly k-1 members of \bar{T}_d .

Then as in the proof of Lemma 2.1, noting that \vec{B}_d contains b-a blocks, that the treatments $(k-1)a+2, \ldots, (k-1)b+1$ do not occur in the set of blocks B_d , and that treatment 1 must be replicated at least once in the set of blocks \vec{B}_d (for otherwise, d is disconnected), one obtains

$$(b-a)k \ge \bar{m}_{1d} + 1 + 2\{(k-1)(b-a) - \bar{m}_{1d}\}$$
,

and

$$\bar{m}_{1d} \le \bar{m}_{2d}(k-1) + (b-a-\bar{m}_{2d})(k-2)$$
,

so that, on simplification, $\bar{m}_{2d} \ge 1$. Thus there is at least one block in \bar{B}_d containing exactly k-1 members of \bar{T}_d and hence of T_d . But this is impossible by the definitions of B_d and \bar{B}_d . This completes the proof.

Lemma 2.3: Suppose $d \in \mathcal{D}$ is not isomorphic to d^* . Then for each $\theta (0 < \theta \le \infty)$,

$$\lambda_{\min}(C_d) < k^{-1} + (1 + k\theta)^{-1}(1 - k^{-1})$$
.

Proof: Since d is not isomorphic to d^* , by Lemma 2.2, without loss of generality, d has two blocks of the form $\{1, 2, ..., k\}, \{\xi, k+1, ..., 2k-1\}$, where $\xi \in \{1, 2, ..., 2k-1\}$, and $r_{id} = 1$ $(2 \le i \le 2k-1)$. Let w be a $v \times 1$ vector defined as

$$w = (-(k-1)(w_1 + w_2), w_1 1'_{k+1}, w_2 1'_{k+1}, 0')'.$$
(2.6)

Then it can be seen that

$$w' R_d w = (k-1)^2 (w_1 + w_2)^2 r_{1d} + (k-1) (w_1^2 + w_2^2) ,$$

$$w' N_d N'_d w = (k-1)^2 \{2 w_2^2 + (w_1 + w_2)^2 (r_{1d} - 1)\} ,$$

$$w' r_d r'_d w = (k-1)^2 (r_{1d} - 1)^2 (w_1 + w_2)^2 ,$$

so that by (1.2), (1.3),

$$w'C_dw - \{k^{-1} + (1+k\theta)^{-1}(1-k^{-1})\}w'w$$

$$= A_{1d}(\theta)(w_1^2 - w_2^2) + A_{2d}(\theta)(w_1 + w_2)^2,$$
(2.7)

where $A_{1d}(\theta) = (k-1)^2 \theta/(1+k\theta)$ and $A_{2d}(\theta)$ depends on θ and the characteristics of the design d but is free from w_1 , w_2 . For $\theta > 0$, $A_{1d}(\theta) > 0$, and it is easy to see that one can choose $(w_1, w_2) \neq (0, 0)$ such that the right-hand side of (2.7) is negative. Since by (2.6), $1'_1w = 0$, the result now follows from (2.7).

From (2.4), (2.5) and Lemma 2.3, it follows that for each θ ($0 < \theta \le \infty$), the designs isomorphic to d^* are uniquely *E*-optimal in \mathcal{D} . For $\theta = 0$, however, uniqueness is not preserved since, as one can easily verify, designs non-isomorphic to d^* can also be *E*-optimal in \mathcal{D} . Anyway, from the above discussion, it is clear that the only designs which are *E*-optimal in \mathcal{D} over the *entire* range $0 \le \theta \le \infty$ are those which are isomorphic to d^* .

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References

- Birkes D, Dodge Y (1991) Optimal $a \times b$ Connected Designs with a+b Observations. J Statist Plann Inf 28:49 59
- Dey A, Bapat RB (1991) Optimal Block Designs with Minimal Number of Observations. Statist ProbLett (to appear)
- Krafft O (1988) On Optimum Designs for Estimating all Contrasts in a Block Model. Preprint
- Mandal NK, Shah KR, Sinha BK (1991) Uncertain Resources and Optimal Designs: Problems and Perspectives. Calcutta Statist Assoc Bull (HK Nandi Memorial Volume) 40:267-282
- Mukerjee R, Chatterjee K, Sen M (1986) D-Optimality of a Class of Saturated Main Effect Plans and Allied Results. Statistics 17:349 355
- Mukerjee R, Sinha BK (1990) Almost Saturated D-Optimal Main Effect Plans and Allied Results. Metrika 37:301 307
- Rao CR (1947) General Methods of Analysis for Incomplete Block Designs. J Amer Statist Assoc 42:541 - 561
- Shah KR, Sinha BK (1989) Theory of Optimal Designs. Lecture Notes in Statistics, No 54. Springer, Berlin