ON HORVITZ-THOMPSON AND DES RAJ ESTIMATORS

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SUMMANY. Under the usual superpopulation set up, it was shown by Go-lambe (1953) that the Morrist-Thompson estimator with a design where the inclusion probabilities are strictly proportional to the a priori expectations will be the best among all the strategies with constant sample sizes, when the super population parameter g is 2. In this paper it is pointed out that when g is different from 2 there exist strategies better than the above strategy.

1. Introduction

Consider a finite population of N units,

We are interested in estimating the total value, Y of a certain characteristic y for this population. The value of y for u_t is denoted by Y_t .

Consider a sample of size two taken without replacement with probabilities of selection being proportional to

$$p_1, p_2, p_2, ..., p_N$$

where

$$\sum_{i=1}^{N} p_i = 1.$$

Let u_i and u_j be the units taken in order of their draw. Des Raj (1956) suggested the following estimate for Y:

$$T_0 = d_1 t_1 + d_2 t_2$$
 ... (1.1)

whore

$$d_1 + d_2 = 1$$
,

and

$$t_1 = \frac{Y_t}{p_t}$$

and

$$l_1 = \frac{Y_i}{p_i}(1-p_i) + Y_i$$

(1.1) being an asymmetric estimator can always be improved by taking the weighted mean of different asymmetric estimators for given unordered sample, the weights being the corresponding conditional probabilities of obtaining the ordered samples given the unordered sample (Halmos, 1949). This estimate happens to be (Murthy, 1957)

$$T_{1} = \frac{1}{2 - p_{i} - p_{j}} \left\{ \frac{Y_{i}}{p_{i}} (1 - p_{j}) + \frac{Y_{j}}{p_{j}} (1 - p_{i}) \right\}. \tag{1.2}$$

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The variance of this estimator is

$$V(T_1) = \sum_{i=1}^{N} \frac{Y_i^2}{p_i} \sum_{j \neq i} p_j \frac{1 - p_i - p_j}{2 - p_i - p_j} - \sum_{j \neq i} Y_i Y_j \left(\frac{1 - p_i - p_j}{2 - p_i - p_i} \right),$$
 ... (1.3)

We shall call the design together with the estimator (1.2) as symmetrised Des Raj strategy.

Consider the sampling design, wherein all the samples are of the same size 2. Let us denote the inclusion probability of the i-th unit by n_i and the joint inclusion probabilities of the i-th and j-th units by n_{ij} . Horvitz and Thompson (1955) suggested the following estimate of the population total Y.

$$T_1 = \frac{Y_i}{\pi_i} + \frac{Y_j}{\pi_i}$$
 ... (1.4)

The variance of this estimate is

$$V(T_2) = \sum_{i=1}^{N} \frac{Y_i^2}{\pi_i} (1 - \pi_i) + \sum_{i \neq j} \sum_{\substack{i \neq j \\ n_i n_j}} \frac{Y_i Y_j}{\pi_i n_j} (\pi_{ij} - \pi_i n_j).$$
 (1.5).

In this paper, we consider designs satisfying

$$\pi_i = 2p_i$$
 ... (1.6)

together with the estimate (1.5), and call it the Horvitz-Thompson strategy.

It is easy to note that there would be situations in which the Horvitz-Thompson strategy is better than symmetrised Des Raj strategy and other situations where symmetrised Des Raj strategy is better, for a fixed value of $(p_1, p_2, ..., p_R)$. So, we would compare the two strategies under a superpopulation set up

2. Superpopulation set up

In many practical situations we would be knowing before hand, the values taken by another characteristic \mathcal{Z} , which is highly correlated with \mathcal{Y} . (We denote the value taken by \mathcal{Z} on u_t by X_t). Now we consider the y-values as the values coming from an infinite population such that the expected value taken by Y_t on u_t for a given value of X_t , is proportional to X_t . (Cochran, 1946). We shall denote the conditional variance of Y_t given X_t by σ_t^2 .

It was observed in field experiments that $V(Y_t|X_t)$ is of the form (Mahalanobis, 1944; Smith, 1938)

$$V(Y_i|X_i) = \sigma^2 X_i^{\theta}$$
, (2.1)

Through intuitive arguments it can be seen that in all practical situations, g lies between 1 and 2. We assume the a priori distribution satisfies (2.2) and (2.4).

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In this paper we will be comparing the two strategies under this model, the criterion for betterness being smaller expected variance, with

$$p_i = \frac{X_i}{Y}. \qquad ... (2.2)$$

Hanurav (1962) has shown that when g=2, the Horvitz-Thompson strategy is better than the symmetrised Des Raj strategy.

3. THE EXPECTED VARIANCES OF THE TWO STRATEGIES

We can write down the expected variances of the two strategies as follows:

$$E_1 = \epsilon(V(T_1)) = \sigma^2 X^g \sum_{i=1}^{N} p_i^{q-1} \sum_{j \neq i} p_j \frac{1 - p_i - p_j}{2 - p_i - p_j}$$
 ... (3.1)

$$E_2 = \varepsilon(V(T_2)) = \sigma^2 X^g \sum p_i^{g-1} \left(\frac{1}{2} - p_i\right)$$
 ... (3.2)

From (3.1) and (3.2), we get

$$E_1 - E_1 = \sigma^2 X^{\theta} \sum_{\ell} p_{\ell}^{\ell-1} \left\{ \sum_{j \neq i} p_j \frac{1 - p_{\ell} - p_j}{2 - p_{\ell} - p_j} - \frac{1}{2} + p_j \right\} = \sigma^2 X^{\theta} \sum_{\ell} p_{\ell}^{\theta-1} a_{\ell} \qquad \dots \quad (3.3)$$

whoro

$$a_i = \frac{1}{2} - \sum_{j \neq i} \frac{p_j}{2 - p_i - p_j}$$
 ... (3.4)

$$= \frac{1}{2} \frac{1}{1-p_i} - \sum_{i} \frac{p_i}{2-p_i-p_i}$$

$$=\frac{1}{2}\sum_{i}\left\{\frac{p_{i}}{1-p_{i}}-\frac{2p_{i}}{2-p_{i}-p_{i}}\right\}$$

$$= \frac{1}{2} \sum_{j} \frac{p_{j}(p_{j} - p_{j})}{(1 - p_{j})(2 - p_{j} - p_{j})} \cdot \dots (3.5)$$

4. MAIN THEOREMS

In the sequel it would be understood that

$$p_1 \leqslant p_2 \leqslant \dots \leqslant p_N. \tag{4.1}$$

Theorem 1: In the usual superpopulation model, the symmetrised Des Raj strategy is superior to the Horvitz-Thompson strategy when g=1 and inferior when g=2, except when all the p's are equal in which case the two strategies coincide.

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Proof: From (3.3) and (3.5), we see that the difference in variance between the symmetrised Des Raj strategy and the Horvitz-Thompson strategy, E, is given by

$$E = \frac{\sigma^2 X^{\epsilon}}{4} \sum_{i=1}^{\infty} \left\{ \frac{p_i^{\epsilon-1} p_i (p_i - p_i)}{(1 - p_i)(2 - p_i - p_j)} + \frac{p_i^{\epsilon-1} p_i (p_j - p_i)}{(1 - p_j)(2 - p_i - p_j)} \right\}$$

$$= \frac{\sigma^2 X^{\epsilon}}{4} \sum_{i=1}^{\infty} \frac{p_i p_i (p_i - p_j)}{(1 - p_j)(1 - p_j)(2 - p_i - p_j)} \left\{ p_i^{\epsilon-2} (1 - p_j) - p_j^{\epsilon-2} (1 - p_i) \right\}$$

$$= \frac{\sigma^2 X^{\epsilon}}{2} \sum_{i=1}^{\infty} \frac{p_i p_j (p_i - p_j)}{(1 - p_j)(2 - p_i - p_j)} \left\{ p_i^{\epsilon-2} (1 - p_j) - p_j^{\epsilon-2} (1 - p_j) \right\} \dots (4.2)$$

$$= \begin{cases} -\frac{\sigma^2 X^{\epsilon}}{2} \sum_{i=1}^{\infty} \sum_{j=1}^{\infty} \frac{(p_i - p_j)^3 (1 - p_i - p_j)}{(1 - p_j)(2 - p_i - p_j)} & \text{if } g = 1 \\ \frac{\sigma^2 X^2}{2} \sum_{i=1}^{\infty} \sum_{j=1}^{\infty} \frac{p_i p_j (p_i - p_j)^3}{(1 - p_j)(1 - p_j)(2 - p_i - p_j)} & \text{if, } g = 2. \dots (4.4) \end{cases}$$

Theorem 1 follows from (4.3) and (4.4)

Lemma : For given values of p1, p2, ..., pN, if the function

$$\sum a_i p_i^{q-1}$$
 ... (4.5)

where

$$a_i = \frac{1}{2} - \sum_{i \neq i} \frac{p_i}{2 - p_i - p_i}$$

and

$$p_i > 0 (\Sigma p_i = 1),$$

is negative when g = g', then the function increases with g at that point.

Proof: We take

$$v_1 \leq v_2 \leq ... \leq v_n$$

It is clear from (3.5) that

$$a_1 \leqslant a_1 \leqslant ... \leqslant a_k \leqslant a_{k+1} \leqslant ... \leqslant a_N$$

where a1, a2, ..., a2 are all negative and a2+1 ... ax are all positive.

Differentiating (4.5) w.r.t. g we get

$$\begin{split} & \Sigma \ a_{l} p_{l}^{l-1} \ \log \ p_{l} = \sum_{l=1}^{k} a_{l} p_{l}^{l-1} \ \log \ p_{l} + \sum_{l=k+1}^{K} a_{l} p_{l}^{l-1} \ \log \ p_{l} \\ \\ & \geqslant \sum_{l=1}^{k} a_{l} p_{l}^{l-1} \ \log \ p_{k} + \sum_{l=k+1}^{K} a_{l} p_{l}^{l-1} \ \log \ p_{k} \\ \\ & \geqslant \left(\sum_{l=1}^{K} a_{l} p_{l}^{l-1}\right) \ \log \ p_{k} \end{split}$$

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where the equality sign holds good only if all the p_i 's are equal. Since $\log p_k$ is negative we get that when g = g'

$$\frac{d}{dq} \sum a_i p_i^{q-1} > 0,$$

unless all the p's are equal. Hence the lemma.

Theorem 2: Under the usual super population model, given $p_1, p_2, ..., p_N$, not all equal, there is a walue for the super population parameter g, say g_0 , where g_0 lies between 1 and 2, such that the Horvitz-Thompson strategy is more precise or less precise (in the expected variance sense) than symmetrised Des Raj's strategy according as $g > g_0$ or $f > g_0$. When $g = g_0$ or $f > g_0$ are all equal the two strategies are equally efficient.

Proof: Theorem follows very easily from Theorem 1 and Lemma 1.

5. NUMERICAL EXAMPLES

We consider here three populations each consisting of 4 units. The values of the characteristic is

population	A	·I	.2	.3	-4	
population	В	.01	.20	.30	.49	
population	С	.23	.24	.26	.27	

The populations were so chosen that in one x's are moderately spread, in one extremely spread and in the last uniform. The officiency of the Horvitz-Thompson strategy compared to the symmetrised Des Raj strategy is tabulated for different values of g.

g	population A	population B	population C
1.0	-9398	.8530	.9988
1-1	.0482	.8752	-9990
1-2	-9568	-8993	.9992
1.3	-9858	.9249	.9994
1.4	.0749	-9513	.0005
1.2	.9843	.9805	.0997
1.0	.9939	1.0054	.0099
1·7	1.0037	1.0326	1.0000
1.8	1.0136	1.0001	1.0003
)·n	1.0238	1.0879	1.0005
2.0	1.0336	1.1470	1.0006

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6. FINAL REMARKS

It was pointed out by Godambe (1955) that the Horvitz-Thompson stategy is the best strategy whenever the above superpopulation model holds good and g takes the value 2. But when g is nearer to 1, evidently there are other estimators which are better. We have to remember that sometimes in practice g comes near to 1 (Fairfield, 1938).

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